

# A New Approach for the Identification of Hidden Markov Models

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# Outline

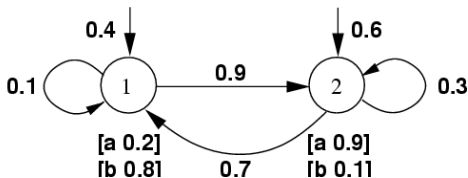
- 1 Problem statement
- 2 Hidden Markov models
- 3 Baum-Welch identification
- 4 Nonnegative matrix factorization
- 5 New identification procedure for HMMs
- 6 Simulation Example
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# Problem statement

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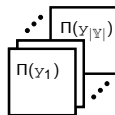
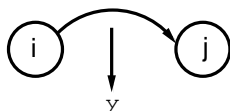
- Approximate identification of hidden Markov models
  - ▶ Baum-Welch
  - ▶ **New technique inspired by subspace identification for LTI systems**

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# Hidden Markov models

- Underlying state process and output process take values from the **finite sets**  $\mathbb{X} = \{1, 2, \dots, |\mathbb{X}|\}$  and  $\mathbb{Y}$  with cardinality  $|\mathbb{Y}|$ 
  - Mealy model  $(\mathbb{X}, \mathbb{Y}, \Pi, \pi(1))$ 
    - $\star \Pi(y)_{ij} = P(x(t+1) = j, y(t) = y | x(t) = i)$
    - $\star \pi(1)_i = P(x(1) = i)$



- Define  $\Pi_{\mathbb{X}} := \sum_{y \in \mathbb{Y}} \Pi(y)$
- Assume stationarity:

$$\pi(1)\Pi_{\mathbb{X}} = \pi(1)$$

# Hidden Markov models

- Define  $\mathbb{Y}^*$  as the set of finite words with symbols from the alphabeth  $\mathbb{Y}$
- **String probabilities**  $\mathcal{P} : \mathbb{Y}^* \mapsto [0, 1]$

$$\mathcal{P}(\mathbf{u}) := P(y(1) = u_1, y(2) = u_2, \dots, y(|y|) = u_{|y|})$$

- String probabilities generated by HMM  $(\mathbb{X}, \mathbb{Y}, \Pi, \pi(1))$

$$\mathcal{P}(\mathbf{u}) = \pi(1)\Pi(\mathbf{u})e$$

where  $\Pi(\mathbf{u}) := \Pi(u_1)\Pi(u_2)\dots\Pi(u_{|\mathbf{u}|})$

- **Identification problem**
  - ▶ *Given:* an output string  $y_1y_2\dots y_T$  of length  $T$
  - ▶ *Find:* HMM  $(\mathbb{X}, \mathbb{Y}, \Pi, \pi(1))$  that models the output string approximately

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# Baum-Welch identification

- Classical approach based on **Maximum Likelihood**
- Solved by **Expectation Maximization**: recursive improvement of the model

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# Nonnegative matrix factorization (1)

- The new identification method makes use of the nonnegative matrix factorization
- **Nonnegative matrix factorization:**  
Given  $M \in \mathbb{R}_+^{m_1 \times m_2}$ , find  $V \in \mathbb{R}_+^{m_1 \times a}$  and  $H \in \mathbb{R}_+^{a \times m_2}$  such that  $M = VH$  with  $a$  as small as possible
- Minimal  $a =$ : positive rank of  $M$
- **Approximate nonnegative matrix factorization** [Lee, Sueng]:  
Given  $M \in \mathbb{R}_+^{m_1 \times m_2}$  and given  $a$ , minimize  $D(M||VH)$  with respect to  $V$  and  $H$ , subject to the constraints  $V, H \geq 0$   
where  $D(A||B)$  is the Kullback-Leibler divergence defined as

$$D(A||B) = \sum_{ij} (A_{ij} \log \frac{A_{ij}}{B_{ij}} - A_{ij} + B_{ij})$$

## Nonnegative matrix factorization (2)

- The algorithm is a special (to preserve nonnegativity) gradient type algorithm

### Theorem (Lee, Sueng)

*The divergence  $D(M||VH)$  is nonincreasing under the update rules*

$$H_{ij} \leftarrow H_{ij} \frac{\sum_{\mu} V_{\mu i} \frac{M_{\mu j}}{(VH)_{\mu j}}}{\sum_{\mu} V_{\mu i}} \quad V_{ki} \leftarrow V_{ki} \frac{\sum_{\nu} H_{i\nu} \frac{M_{k\nu}}{(VH)_{k\nu}}}{\sum_{\nu} H_{i\nu}}$$

*$(V, H)$  is invariant under these updates if and only if  $(V, H)$  is a stationary point of the divergence.*

- Applications: e.g. in image representations

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# New identification procedure for HMMs

Idea behind "subspace identification":

- 1 Determine state sequence directly from data
- 2 Determine system parameters from state/output sequence

Proposed HMM identification method:

- 1 Estimate of the **state sequence**  
→ from the given output string
- 2 Estimate of the **system matrices**  
→ from the obtained state sequence and the given output sequence

# Estimating the state sequence (1)

- Given the model, determination of state sequence corresponding to a particular output sequence
- Suppose we are given:
  - the matrix  $M(i_1, i_2)$  of the underlying model  $(\mathbb{X}, \mathbb{Y}, \Pi, \pi(1))$

$$M(i_1, i_2)_{kl} = \mathcal{P}(\mathbf{u}_k \mathbf{v}_l)$$

with  $\mathbf{u}_k$  the  $k$ -th length- $i_1$  string and  $\mathbf{v}_l$  the  $l$ -th length  $i_2$  string

$$M(i_1, i_2) = \begin{bmatrix} \mathcal{P}(\mathbf{u}_1 \mathbf{v}_1) & \mathcal{P}(\mathbf{u}_1 \mathbf{v}_2) & \dots & \mathcal{P}(\mathbf{u}_1 \mathbf{v}_{|\mathbb{Y}|i_2}) \\ \mathcal{P}(\mathbf{u}_2 \mathbf{v}_1) & \mathcal{P}(\mathbf{u}_2 \mathbf{v}_2) & \dots & \mathcal{P}(\mathbf{u}_2 \mathbf{v}_{|\mathbb{Y}|i_2}) \\ \vdots & \vdots & \ddots & \vdots \\ \mathcal{P}(\mathbf{u}_{|\mathbb{Y}|i_2} \mathbf{v}_1) & \mathcal{P}(\mathbf{u}_{|\mathbb{Y}|i_2} \mathbf{v}_2) & \dots & \mathcal{P}(\mathbf{u}_{|\mathbb{Y}|i_2} \mathbf{v}_{|\mathbb{Y}|i_2}) \end{bmatrix}$$

## Estimating the state sequence (2)

- ▶ the nonnegative decomposition of  $M(i_1, i_2)$  given by

$$M(i_1, i_2) = VH \quad (1)$$

with

$$V = \begin{bmatrix} \pi(1)\Pi(\mathbf{u}_1) \\ \pi(1)\Pi(\mathbf{u}_2) \\ \vdots \\ \pi(1)\Pi(\mathbf{u}_{|\mathbb{Y}|i_1}) \end{bmatrix}$$
$$H = \begin{bmatrix} \Pi(\mathbf{v}_1)\mathbf{e} & \Pi(\mathbf{v}_2)\mathbf{e} & \dots & \Pi(\mathbf{v}_{|\mathbb{Y}|i_2})\mathbf{e} \end{bmatrix}$$

- ▶ Rows of  $V$  can be seen to be

$$V_{k,i} = P(y(1,2,\dots,i_1) = \mathbf{u}_k, x(i_1 + 1) = i)$$



## Estimating the state sequence (3)

- Using  $V$ , compute  $P(x(t)|y(t-i_1)\dots y(t-1))$  for  $t = i_1 + 1, i_1 + 2, \dots, T$
- Now replace the probabilities by a maximum likelihood estimate  $\hat{x}(t) \in \{0, 1\}^{1 \times |\mathbb{X}|}$

$$\hat{x}(t)_i = 1$$

$$\hat{x}(t)_j = 0, \quad j = 1, \dots, i-1, i+1, \dots, |\mathbb{X}|$$

where  $i$  is the most likely state estimate for  $x(t)$  given the observations of  $y(t-i_1)\dots y(t-1)$

- Define the estimated state sequence matrix  $\hat{X}_{i_1} \in \{0, 1\}^{(T-i_1) \times |\mathbb{X}|}$  as

$$\hat{X}_{i_1} = \begin{bmatrix} \hat{x}(i_1 + 1) \\ \hat{x}(i_1 + 2) \\ \vdots \\ \hat{x}(T) \end{bmatrix}$$

## Estimating the state sequence (4)

- The next step is to estimate the state at time  $t + 1$  based on  $y(t - i_1) \dots y(t - 1)$
- This is done in a similar way using  $M(i_1 + 1, i_2)$
- We obtain

$$\hat{X}_{i_1+1}^+ = \begin{bmatrix} \hat{x}^+(i_1 + 2) \\ \hat{x}^+(i_1 + 3) \\ \vdots \\ \hat{x}^+(T + 1) \end{bmatrix}.$$

# Estimating the string probabilities

- The matrices  $M(i_1, i_2)$  and  $M(i_1 + 1, i_2)$  contain string probabilities of length  $i_1 + i_2$  and  $i_1 + i_2 + 1$
- String probabilities can be estimated from the given output sequence yielding  $\hat{M}(i_1, i_2)$  and  $\hat{M}(i_1 + 1, i_2)$
- The nonnegative matrix factorization yields

$$\begin{aligned}\hat{M}(i_1, i_2) &= \hat{V}\hat{H} \\ \hat{M}(i_1 + 1, i_2) &= \hat{W}\hat{H}\end{aligned}$$

## Estimating the system matrices

- We now calculate the system matrices using

$$P(x(t+1), y(t) = \mathcal{Y} \mid y(t-i_1, \dots, t-1) = u_{t-i_1} \dots u_{t-1}) = \\ P(x(t) \mid y(t-i_1, \dots, t-1) = u_{t-i_1} \dots u_{t-1}) \Pi(\mathcal{Y}),$$

- Define

$$\hat{x}^{\mathcal{Y}}(t+1) = \begin{cases} \hat{x}^+(t+1) & u_t = \mathcal{Y} \\ [0 \ 0 \ \dots \ 0] & \text{else} \end{cases}$$

- Now  $\hat{\Pi}(\mathcal{Y}), \mathcal{Y} \in \mathbb{Y}$  can be calculated from the least squares problem

$$\begin{bmatrix} \hat{X}_{i_1+1}^{\mathcal{Y}_1} & \hat{X}_{i_1+1}^{\mathcal{Y}_2} & \dots & \hat{X}_{i_1+1}^{\mathcal{Y}_{|\mathbb{Y}|}} \end{bmatrix} = \hat{X}_{i_1} \begin{bmatrix} \hat{\Pi}(\mathcal{Y}_1) & \hat{\Pi}(\mathcal{Y}_2) & \dots & \hat{\Pi}(\mathcal{Y}_{|\mathbb{Y}|}) \end{bmatrix}$$

- It can be proven that the matrices  $\hat{\Pi}(\mathcal{Y}), \mathcal{Y} \in \mathbb{Y}$  are elementwise nonnegative (due to ML approach)
- The initial state distribution  $\hat{\pi}(1)$  is taken equal to the normalised left eigenvector of  $\sum_{\mathcal{Y}} \hat{\Pi}(\mathcal{Y})$  corresponding to the eigenvalue 1

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# Simulation Example (1)

- Given an output string  $y_1 \dots y_{1000}$  generated by  $\lambda_{\text{true}} = (\{1,2\}, \{1,2\}, \Pi_{\text{true}}, \pi_{\text{true}}(1))$  where

$$\Pi_{\text{true}}(1) = \begin{bmatrix} 0.20 & 0.40 \\ 0.00 & 0.20 \end{bmatrix}$$

$$\Pi_{\text{true}}(2) = \begin{bmatrix} 0.10 & 0.30 \\ 0.80 & 0.00 \end{bmatrix}$$

$$\pi_{\text{true}} = [ 0.53 \quad 0.47 ]$$

- Identification ( $|\mathbb{X}_{\text{SS}}| = 2$ ) with new method with  $i_1 = i_2 = 3$

$$\Pi_{\text{SS}}(1) = \begin{bmatrix} 0.0699 & 0.2574 \\ 0.5651 & 0.0000 \end{bmatrix}$$

$$\Pi_{\text{SS}}(2) = \begin{bmatrix} 0.1342 & 0.5386 \\ 0.4349 & 0.0000 \end{bmatrix}$$

$$\pi_{\text{SS}} = [ 0.5568 \quad 0.4432 ]$$

- Identification ( $|\mathbb{X}_{\text{BW}}| = 2$ ) with Baum-Welch

$$\Pi_{\text{BW}}(1) = \begin{bmatrix} 0.0736 & 0.0986 \\ 0.5311 & 0.1415 \end{bmatrix}$$

$$\Pi_{\text{BW}}(2) = \begin{bmatrix} 0.0751 & 0.7526 \\ 0.2424 & 0.0850 \end{bmatrix}$$

$$\pi_{\text{BW}} = [ 0 \quad 1 ]$$

## Simulation Example (2)

- Distance between HMMs

$$D(\lambda_{\text{true}} || \lambda_{\text{approx}}) = \sum_{\mathbf{u} \in \mathbb{Y}^*} \mathcal{P}(\mathbf{u} | \lambda_{\text{true}}) \log \frac{\mathcal{P}(\mathbf{u} | \lambda_{\text{true}})}{\mathcal{P}(\mathbf{u} | \lambda_{\text{approx}})}$$

- In practice

$$D_{\Delta}(\lambda_{\text{true}} || \lambda_{\text{approx}}) = \sum_{\mathbf{u} \in \mathbb{Y}^*, |\mathbf{u}| \leq \Delta} \mathcal{P}(\mathbf{u} | \lambda_{\text{true}}) \log \frac{\mathcal{P}(\mathbf{u} | \lambda_{\text{true}})}{\mathcal{P}(\mathbf{u} | \lambda_{\text{approx}})}$$

- Results for  $\Delta = 8$

$$D_{\Delta}(\lambda_{\text{true}} || \lambda_{\text{SS}}) = 0.3876$$

$$D_{\Delta}(\lambda_{\text{true}} || \lambda_{\text{BW}}) = 1.8955$$

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# Conclusions

- New identification procedure for hidden Markov models in two steps
  - ▶ Calculate the state sequence directly from output sequence
  - ▶ Estimate system matrices from output sequence and estimated state sequence
  
- Method gives good results on a simulation example