ARMAX System Identification: First X, then AR, finally MA

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(joint work with Ivan Markovsky and Bart L.M. De Moor)

In this extended abstract, 'process' means: a zero mean, gaussian, stationary, ergodic vector process on \mathbb{Z} , \bot means 'independence', and 'white noise' means a process ε for which the $\sigma^t \varepsilon(0)$'s are all \bot for $t \in \mathbb{Z}$, and σ denotes the shift $(\sigma f(t) := f(t+1))$. Consider the difference equation

$$W(\sigma)w = E(\sigma)\varepsilon,$$
 (ARMAX)

with W, E suitably sized polynomial matrices. The *behavior* of (ARMAX) consists of all processes w such that (ARMAX) holds for some white noise process ε . The identification (ID) problem is to obtain estimates of (W, E) from observation of a realization of w:

$$ilde{w}(1), ilde{w}(2), \dots, ilde{w}(T).$$

In this extended abstract, we will assume for simplicity of exposition that $T = \infty$. In the actual algorithm, we assume T finite, and study the behavior of the estimates as $T \to \infty$.

Every ARMAX system admits a more refined representation

$$A(\sigma)R(\sigma)w = M(\sigma)\varepsilon$$
 (AR-MA-X)

with A square, det(A) non-zero and without unit circle roots, and R left-prime. Note that $R(\sigma)w = 0$ corresponds to the 'exogenous' part of the AR-MA-X system (obtained by setting $\varepsilon = 0$). We call R the 'X' (exogenous) part, A the 'AR' part,

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and M the 'MA' part of the AR-MA-X system. We present an algorithm that identifies first R, then A, and finally M.

Many interesting problems emerge: When do two systems (A, R, M) define the same behavior? Obtain canonical forms. If $w = \begin{bmatrix} u \\ y \end{bmatrix}$, when is u a 'free input', in the sense that for any process u, there exists a process y such that $w = \begin{bmatrix} u \\ y \end{bmatrix}$ belongs to the behavior of (AR-MA-X)? When is this y unique? In [1] these issues are studied in depth.

It is easy to see that for all $n \in \mathbb{R}[\xi]$ in the $\mathbb{R}[\xi]$ -module generated by the transposes of the rows of R, $n(\sigma)^{\top}w \perp \varepsilon$. Assume that $R = [P \ Q]$ with P square, and correspondingly $w = \begin{bmatrix} u \\ y \end{bmatrix}$, with $u \perp \varepsilon$. Now look for the finite linear combinations of the rows of the observed

$$\tilde{W} = \begin{bmatrix} \tilde{w}(1) & \tilde{w}(2) & \tilde{w}(3) & \cdots & \tilde{w}(t) & \cdots \\ \tilde{w}(2) & \tilde{w}(3) & \tilde{w}(4) & \cdots & \tilde{w}(t+1) & \cdots \\ \tilde{w}(3) & \tilde{w}(4) & \tilde{w}(5) & \cdots & \tilde{w}(t+2) & \cdots \\ \vdots & \vdots & \vdots & \vdots & \vdots & \ddots \end{bmatrix}$$

that are orthogonal to the rows of the observed

$$\tilde{U} = \begin{bmatrix} \tilde{u}(1) & \tilde{u}(2) & \tilde{u}(3) & \cdots & \tilde{u}(t) & \cdots \\ \tilde{u}(2) & \tilde{u}(3) & \tilde{u}(4) & \cdots & \tilde{u}(t+1) & \cdots \\ \tilde{u}(3) & \tilde{u}(4) & \tilde{u}(5) & \cdots & \tilde{u}(t+2) & \cdots \\ \vdots & \vdots & \vdots & \vdots & \vdots & \ddots \end{bmatrix}$$

Call these linear combinations 'orthogonalizers'. Obviously each orthogonalizer is a vector of the form $\pi = \operatorname{col}(\pi_0, \pi_1, \dots, \pi_n, \dots)$, with the π_n 's $\in \mathbb{R}^w$, and all but a finite number of them non-zero. Organize the orthogonalizers as polynomial vectors $\pi(\xi) = \pi_0 + \pi_1 \xi + \dots + \pi_n \xi^n + \dots \in \mathbb{R}^w[\xi]$.

It can be shown that if \tilde{u} is persistently exciting, then the orthogonalizers form exactly the $\mathbb{R}[\xi]$ -module generated by the transposes of the rows of R. This yields an algorithm for identifying R from the observations via the orthogonalizers. As we have described it here, this algorithm requires an infinite number of rows of \tilde{W} and \tilde{U} , but if we assume that (upper bounds for) the lag L and the dynamic order n of the AR-MA-X system are known, we can restrict attention to the first L rows of \tilde{W} and the first L + n rows of \tilde{U} .

Once *R* has been estimated, we compute

$$\tilde{a}=\hat{R}(\boldsymbol{\sigma})\tilde{w},$$

and obtain an estimate \hat{A} of A from \tilde{a} , and proceed by computing

$$\tilde{m}=\hat{A}(\boldsymbol{\sigma})\tilde{a},$$

to obtain an estimate \hat{M} of M, leading to an estimate $(\hat{R}, \hat{A}, \hat{M})$ for (R, A, M).

This extended abstract reports on research in progress. A full paper is in preparation.

References

[1] E.J. Hannan and M. Deistler, *The Statistical Theory of Linear Systems*, Academic Press, 1979.